

Dealing with Heteroskedasticity

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Biol 520C: Statistical modelling for biological data

1. Heteroskedasticity and the IID Assumption
2. Variance Structures

Heteroskedasticity and the IID Assumption

We started with simple linear regression:

$$y_i = \beta_0 + \beta x_i + \varepsilon_i$$

We then extended this to multiple linear regression of the form:

$$y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \dots + \beta_n x_{ni} + \varepsilon_i$$

And then to linear mixed effects model that account for correlations within groups:

$$\mathbf{y}_i = \mathbf{X}_i \boldsymbol{\beta} + \mathbf{Z}_i \mathbf{b}_i + \varepsilon_i$$

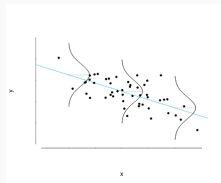
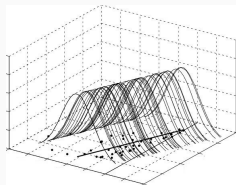
But the elephant in the room is our assumption that $\varepsilon_i \sim \mathcal{N}(0, \sigma^2)$

The models we've been working with so far assume that the residuals at each x_i value should be Independent and Identically Distributed (IID).

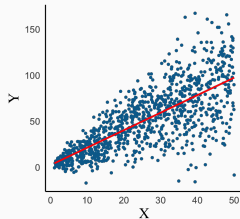
More specifically, they should be normally distributed with a constant variance σ^2 :

$$\varepsilon_i \sim \mathcal{N}(0, \sigma^2)$$

In real systems variances can differ across groups, times, etc. (i.e., heteroskedasticity) and the 'identical' part of the IID assumption is likely to be broken in many cases.



Wikipedia: In statistics, a vector of random variables is heteroskedastic (from Ancient Greek hetero 'different' and skedasis 'dispersion') if the variability of the random disturbance is different across elements of the vector.



Formally, for a hypothetical dataset with three observations, A is homoskedastic; in B, the variance increases steadily across samples; in C, the variance depends on the value of x .

$$A = \sigma^2 \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix} \quad B = \sigma^2 \begin{bmatrix} 1 & 0 & 0 \\ 0 & 2 & 0 \\ 0 & 0 & 3 \end{bmatrix} \quad C = \sigma^2 \begin{bmatrix} x_1 & 0 & 0 \\ 0 & x_2 & 0 \\ 0 & 0 & x_3 \end{bmatrix}$$

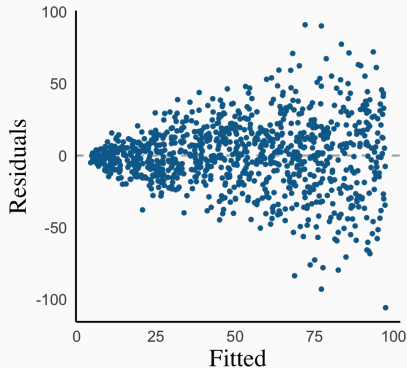
More generally, if the variance-covariance matrix of ε_i across i has a non-constant diagonal, the errors will be heteroskedastic.

What heteroskedasticity boils down to is non-constant variance that varies with other components of the system.

When the IID assumption is violated by heteroskedasticity variances, and hence standard errors, will be poorly estimated resulting in misleading p -values (we'll explore this concept in this week's practical).

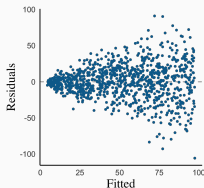
The deterministic part of the model *may* be correctly specified, but any predictions are likely to be off.

The easiest way to identify the presence of heteroskedasticity is by plotting a model's residuals against the predicted values.



These cone shaped residuals are the tell-tale sign of heteroskedasticity.

Identify the presence of heteroskedasticity is fairly easy.



In order to deal with the issue we need to identify the cause (i.e., what group/factor does the variance change with).

The easiest way to identify the cause of heteroskedasticity is by plotting a model's residuals against each of the predictors.

Today we're going to work with a dataset collected by Smith *et al.* (2005) to examine seasonal patterns in reproductive and somatic tissues in the squid *Loligo forbesi*.

The data are comprised of measures of testis weight (TW) and dorsal mantle length (DML) from 768 male squid. The month of data collection was also recorded.

Today's starting point is the linear regression model:

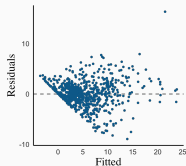
$$TW_i = \beta_0 + \beta_1 DML_i + \beta_2 \text{month}_i + \beta_3 DML_i : \text{month}_i + \varepsilon_i$$

```
library(nlme)

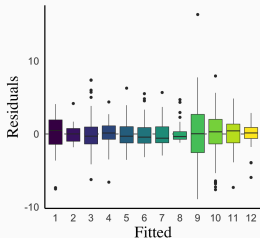
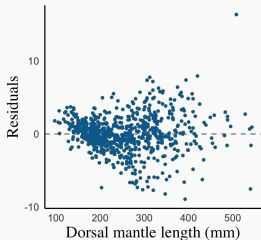
data <- read.csv("Squid.csv")

data$month <- as.factor(data$month)

mod <- gls(TW ~ DML*month, data = data)
```

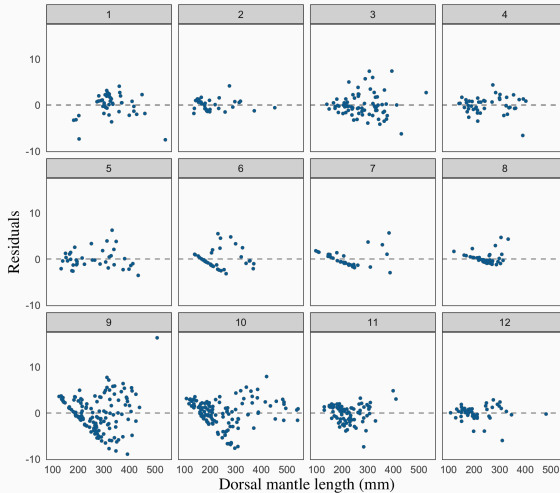


These data are clearly heteroskedastic. To identify the cause we'll look at the model's residuals against DML and month.



$\sigma^2 \propto \text{DML}$, but different months also have different σ^2 .

Residual spread changes month to month, and in some months increases with DML. So both need to be modelled for a full correction.



summary(mod)

	Value	Std. Error	t-value	p-value
(Intercept)	3.215222	1.8740686	1.715637	0.0866
DML	0.021157	0.0055585	3.806244	0.0002
month2	-6.720020	2.3282175	-2.886337	0.0040
month3	-3.926923	2.2081953	-1.778340	0.0758
month4	-4.772045	2.2875351	-2.086108	0.0373
month5	-2.771457	2.2656715	-1.223239	0.2216
month6	-9.598061	2.4163109	-3.972196	0.0001
month7	-7.494959	2.2933137	-3.268178	0.0011
month8	-7.479426	2.9310875	-2.551758	0.0109
month9	-14.963009	2.0353133	-7.351698	0.0000
month10	-12.320827	1.9637718	-6.274063	0.0000
month11	-12.650490	2.1922272	-5.770611	0.0000
month12	-9.235813	2.2264505	-4.148223	0.0000
DML:month2	0.018032	0.0083101	2.169894	0.0303
DML:month3	0.003151	0.0068657	0.458989	0.6464
DML:month4	0.002972	0.0074316	0.399848	0.6894
DML:month5	-0.008677	0.0072378	-1.198801	0.2310
DML:month6	0.017620	0.0084180	2.093183	0.0367
DML:month7	0.004647	0.0079929	0.581432	0.5611
DML:month8	0.000500	0.0105569	0.047373	0.9622
DML:month9	0.044242	0.0062130	7.120835	0.0000
DML:month10	0.039495	0.0059673	6.618449	0.0000
DML:month11	0.046671	0.0074640	6.252829	0.0000
DML:month12	0.034099	0.0077022	4.427127	0.0000

Running summary on the fit shows a lot of significant parameters.

But because of the heteroskedasticity we can't trust the p -values.

We could try to transform the data, but the heteroskedasticity is interesting ecological and we should avoid treating it as a 'nuisance' and try to model it.

Variance Structures

One of the simplest variance structures is the so called 'fixed variance structure'.

$$\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times x_i)$$

This formulation allows for variance to increase with x

Because σ^2 is already being estimated, and x_i are our data, there are no extra parameters to fit with this formulation.

In other words, we pick this structure up for free, which is big bonus, but it also means we don't learn anything new about the system.

For our squid data, variance increased with DML, so a reasonable formulation to handle the heteroskedasticity would be:

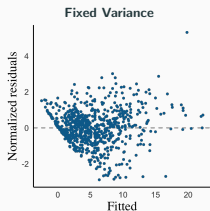
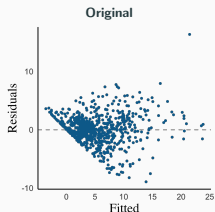
$$\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times \text{DML}_i)$$

We can do this via the `weights` argument in the `glS()` function.

```
mod2 <- gls(TW ~ DML*month,
            weights = varFixed(~DML),
            data = data)

anova(mod, mod2)
```

	Model	df	AIC	BIC	logLik
mod	1	25	3752.084	3867.385	-1851.042
mod2	2	25	3620.898	3736.199	-1785.449



AIC shows an improvement, but the residuals still show signs of heterosk.

Another way for heteroskedasticity to enter into the system is for different **categorical** groups to have different variances (e.g., males/females, seasons, years, etc...).

$$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma_j^2)$$

If there is only one group, the variance is constant. When there is a grouping factor with $j > 1$ levels, the variance function allows j different variances, one for each level of the factor.

In practice, the variance function represents the ratios between the different σ^2 and a reference σ^2 , so $j - 1$ new coefficients are needed to represent the variance function.

$$\sigma_1^2 = \sigma^2 \times 1, \quad \sigma_2^2 = \sigma^2 \times \theta_1 \quad \dots \quad \sigma_j^2 = \sigma^2 \times \theta_{j-1}$$

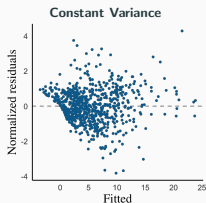
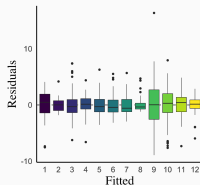
For our squid data, variance differed across months, so we can have the variance change across months: $\varepsilon_{i,month} \sim \mathcal{N}(0, \sigma_{month}^2)$

We can do this via the `varIdent()` function.

```
mod3 <- gls(TW ~ DML*month,
            weights = varIdent(form = ~1|month),
            data = data)

summary(mod3)
...
Variance function:
Structure: Different standard deviations per stratum
Formula: ~1 | month
Parameter estimates:
           2           9           12           11           8
1.0000000 2.9913585 1.2736165 1.5090469 0.9821378
           10           5           7           6           4
2.2162169 1.6396315 1.3783514 1.6473098 1.4232364
           1           3
1.9584902 1.9788664
...
anova(mod, mod3)

      Model df      AIC      BIC    logLik    Test  L.Ratio p-value
mod      1 25 3752.084 3867.385 -1851.042
mod3     2 36 3614.436 3780.469 -1771.218 1 vs 2 159.6479 <.0001
```



Again, AIC shows an improvement, but the residuals aren't convincing.

The third variance structure we'll look at is the power variance structure:

$$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 \times |x_i|^{2\delta_j})$$

The power variance structure is a generalisation of the fixed variance structure.

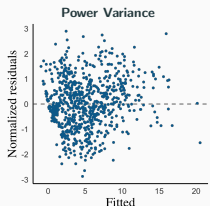
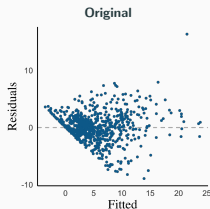
- If $\delta = 0$ this reduces to the linear regression model.
- If $\delta = \frac{1}{2}$ this reduces to the fixed variance structure.
- If $\delta > \frac{1}{2}$ the variance is allowed to spread more than a fixed variance for the same increase in x .
- When a grouping factor is present, a different δ is used for each factor level.

For our squid data, variance increased with DML, so a reasonable formulation would be: $\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times |DML_i|^{2\delta})$

We can do this via the `varPower()` function.

```
mod4 <- gls(TW ~ DML*month,
            weights = varPower(form = ~ DML),
            data = data)
summary(mod4)
...
Variance function:
Structure: Power of variance covariate
Formula: ~DML
Parameter estimates:
    power
1.759009
...
anova(mod, mod4)

      Model df      AIC      BIC    logLik  Test  L.Ratio p-value
mod      1  25 3752.084 3867.385 -1851.042
mod4     2  26 3473.019 3592.932 -1710.509 1 vs 2 281.0648 <.0001
```



AIC shows a big improvement, and the residuals look much better.

Variance increased with DML and across months, so we could also try

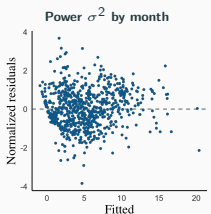
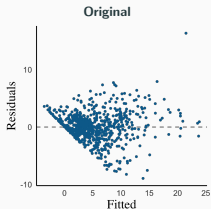
$$\epsilon_{i,\text{month}} \sim \mathcal{N}(0, \sigma^2 \times |DML_i|^{2\delta_{\text{month}}})$$

```
mod5 <- gls(TW ~ DML*month,
            weights = varPower(form = ~ DML|month),
            data = data)

summary(mod5)

...
Variance function:
Structure: Power of variance covariate, different strata
Formula: ~DML | month
Parameter estimates:
      2      9      12      11      8      10
1.728531 1.789503 1.733561 1.749264 1.617694 1.789187
      5      7      6      4      1      3
1.746517 1.673244 1.754488 1.711371 1.698945 1.722742
...
anova(mod, mod5)

      Model df      AIC      BIC    logLik  Test L.Ratio p-value
mod      1 25 3752.084 3867.385 -1851.042
mod5     2 37 3407.511 3578.156 -1666.755 1 vs 2 368.5728 <.0001
```



Best AIC yet, and the residuals look pretty good.

The next variance structure we'll look at is the exponential variance structure:

$$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 \times e^{2\delta_j \times x_i})$$

The exponential variance structure models the variance of the residuals as σ^2 multiplied by an exponential function.

- If $\delta = 0$ this reduces to the linear regression model.
- If $\delta > 0$ the variance increases with x .
- If $\delta < 0$ the variance decreases with x .
- As with the power variance, δ can be grouped by a categorical variable j .

For our squid data, variance increased with DML, so our formulation is:

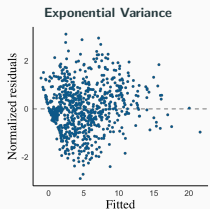
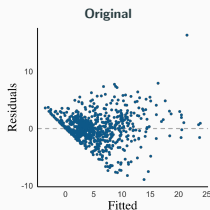
$$\epsilon_i \sim \mathcal{N}(0, \sigma^2 \times e^{2\delta \times DML_i})$$

We can implement this via the `varExp()` function.

```
mod6 <- gls(TW ~ DML*month,
            weights = varExp(form = ~ DML),
            data = data)

summary(mod6)
...
Variance function:
Structure: Exponential of variance covariate
Formula: ~DML
Parameter estimates:
    expon
0.006832722
...
anova(mod, mod6)

      Model df      AIC      BIC    logLik  Test  L.Ratio p-value
mod      1  25 3752.084 3867.385 -1851.042
mod6     2  26 3478.152 3598.066 -1713.076 1 vs 2 275.9312 <.0001
```

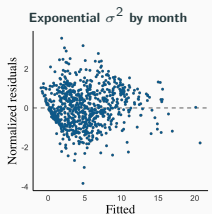
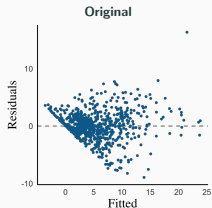


AIC shows an improvement, and the residuals look ok.

Again, we can assign different variances for different months via the `varExp()` function.

```
mod7 <- gls(TW ~ DML*month,
            weights = varExp(form = ~ DML|month),
            data = data)
summary(mod7)
...
Variance function:
Structure: Exponential of variance covariate, different strata
Formula: ~DML | month
Parameter estimates:
      2      9      12      11      8
0.005463411 0.007471853 0.006089471 0.006700889 0.004340133
      10      5      7      6      4
0.007371252 0.006060562 0.005575366 0.006959413 0.005612065
      1      3
0.005510245 0.006214627
...
anova(mod, mod7)

      Model df      AIC      BIC    logLik    Test  L.Ratio p-value
mod      1 25 3752.084 3867.385 -1851.042
mod7     2 37 3419.719 3590.364 -1672.859 1 vs 2 356.3652 <.0001
```



AIC shows an improvement, and the residuals look good.

The penultimate variance structure we'll look at is the constant plus power variance structure:

$$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 \times (\delta_{1j} + |x_i|^{\delta_{2j}})^2)$$

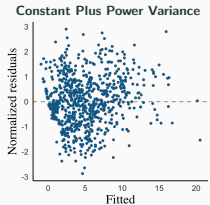
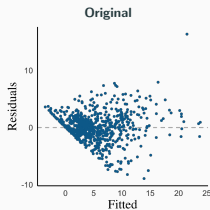
- If $\delta_1 = 0$ & $\delta_2 = 0$ this reduces to the linear regression model.
- If $\delta_1 = 0$ & $\delta_2 \neq 0$ this reduces to the power variance.
- Here δ_1 & δ_2 can be grouped by a categorical variable j .

For our squid data our formulation is: $\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times (\delta_1 + |DML_i| \delta_2)^2)$

Which we implement via `varConstPower()`.

```
mod8 <- gls(TW ~ DML*month,
            weights = varConstPower(form = ~ DML),
            data = data)
summary(mod8)
...
Variance function:
Structure: Constant plus power of variance covariate
Formula: ~DML
Parameter estimates:
      const      power
0.0992986  1.7590192
...
anova(mod, mod8)

      Model df      AIC      BIC    logLik  Test L.Ratio p-value
mod      1  25 3752.084 3867.385 -1851.042
mod8     2  27 3475.019 3599.544 -1710.509 1 vs 2 281.065 <.0001
```

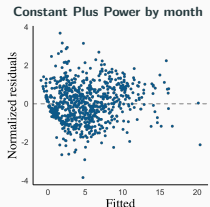
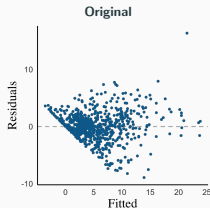


AIC shows an improvement, and the residuals look ok.

Again, we can assign different variances for different months via the `varConstPower()` function.

```
mod9 <- gls(TW ~ DML*month,
            weights = varConstPower(form = ~ DML|month),
            data = data)
summary(mod9)
...
Variance function:
Structure: Constant plus power of variance covariate, different
strata
Formula: ~DML | month
Parameter estimates:
           2           9           12           11           8
const 0.1098482 0.1044132 0.0978525 0.09883092 0.09952452
power 1.7285761 1.7895481 1.7336082 1.74931089 1.61774116
           10           5           7           6           4
const 0.09361078 0.09794217 0.09729915 0.09819052 0.0975102
power 1.78923387 1.74656454 1.67329035 1.75453620 1.7114162
           1           3
const 0.0988839 0.09964889
power 1.6989915 1.72278732
...
anova(mod, mod9)

      Model df      AIC      BIC    logLik    Test  L.Ratio p-value
mod      1 25 3752.084 3867.385 -1851.042
mod9     2 49 3431.511 3657.501 -1666.755 1 vs 2 368.5728 <.0001
```



AIC and residuals look good, but note all the DFs!

The last variance structure we'll look at is the combined variance structure.

This is an R function (`varComb`) as opposed to a specific formulation, and allows you to flexibly combine the previously described structures.

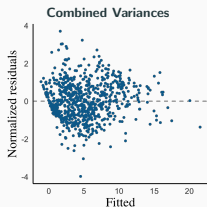
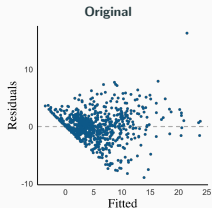
The corresponding variance function is equal to the product of the variance functions.

It's very flexible, but challenging to work with.

```

mod10 <- gls(TW ~ DML*month,
  weights = varComb(varIdent(form = ~1|month),
    varExp(form = ~DML)),
  data = data)
summary(mod10)
...
Combination of variance functions:
Structure: Different standard deviations per stratum
Formula: ~1 | month
Parameter estimates:
      2          9          12          11          8          10
1.0000000 1.4806117 1.0577473 1.1948695 0.5829504 1.4255755
      5          7          6          4          1          3
1.1276263 0.7353367 1.2712401 0.9710625 0.8668377 1.0561607
Structure: Exponential of variance covariate
Formula: ~DML
Parameter estimates:
  expon
0.006711743
...
anova(mod, mod10)

      Model df      AIC      BIC    logLik    Test  L.Ratio p-value
mod      1 25 3752.084 3867.385 -1851.042
mod10    2 37 3414.817 3585.463 -1670.409 1 vs 2 361.2663 <.0001
  
```



We just fit 10 different models, but how do we know which variance structure to go with?

Not all variance structures are nested, so the easiest way to pick the best one is by AIC based model selection.

```
#Calculate AIC values
TABLE <- AIC(mod, mod2, mod3, mod4,
             mod5, mod6, mod7, mod8,
             mod9, mod10)

#Ordered by lowest to highest AIC
TABLE <- TABLE[order(TABLE$AIC),]

#Calculate Delta AICs
TABLE$DeltaAIC <- TABLE$AIC - TABLE$AIC[1]

#Evidence compared to AIC best model
TABLE$Evidence <- 1/exp(-(TABLE$DeltaAIC/2))
```

	df	AIC	DeltaAIC	Evidence
mod5	37	3407.511	0.000000	1.000000e+00
mod10	37	3414.817	7.306579	3.860144e+01
mod7	37	3419.719	12.207673	4.475717e+02
mod9	49	3431.511	24.000039	1.627580e+05
mod4	26	3473.019	65.507991	1.678338e+14
mod8	27	3475.019	67.507874	4.561929e+14
mod6	26	3478.152	70.641592	2.185883e+15
mod3	36	3614.436	206.924992	8.574149e+44
mod2	25	3620.898	213.386607	2.169267e+46
mod	25	3752.084	344.572838	6.653361e+74

Original Model

	Value	Std. Error	t-value	p-value
(Intercept)	3.215222	1.8740686	1.715637	0.0866
DML	0.021157	0.0055585	3.806244	0.0002
month2	-6.720020	2.3282175	-2.886337	0.0040
month3	-3.926923	2.2081953	-1.778340	0.0758
month4	-4.772045	2.2875351	-2.086108	0.0373
month5	-2.771457	2.2656715	-1.223239	0.2216
month6	-9.598061	2.4163109	-3.972196	0.0001
month7	-7.494959	2.2933137	-3.268178	0.0011
month8	-7.479426	2.9310875	-2.551758	0.0109
month9	-14.963009	2.0353133	-7.351698	0.0000
month10	-12.320827	1.9637718	-6.274063	0.0000
month11	-12.650490	2.1922272	-5.770611	0.0000
month12	-9.235813	2.2264505	-4.148223	0.0000
DML: month2	0.018032	0.0083101	2.169894	0.0303
DML: month3	0.003151	0.0068657	0.458989	0.6464
DML: month4	0.002972	0.0074316	0.399848	0.6894
DML: month5	-0.008677	0.0072378	-1.198801	0.2310
DML: month6	0.017620	0.0084180	2.093183	0.0367
DML: month7	0.004647	0.0079929	0.581432	0.5611
DML: month8	0.000500	0.0105569	0.047373	0.9622
DML: month9	0.044242	0.0062130	7.120835	0.0000
DML: month10	0.039495	0.0059673	6.618449	0.0000
DML: month11	0.046671	0.0074640	6.252829	0.0000
DML: month12	0.034099	0.0077022	4.427127	0.0000

Selected Model

	Value	Std. Error	t-value	p-value
(Intercept)	-4.780799	1.3347084	-3.581905	0.0004
DML	0.046587	0.0047135	9.883629	0.0000
month2	1.334866	1.6540452	0.807031	0.4199
month3	3.661194	1.5402841	2.376960	0.0177
month4	3.410094	1.5457315	2.206136	0.0277
month5	3.953020	1.5990089	2.472169	0.0137
month6	-0.666428	1.7407663	-0.382836	0.7020
month7	3.669079	1.3886026	2.642281	0.0084
month8	3.241259	1.4361065	2.256977	0.0243
month9	-1.646933	1.4697410	-1.120560	0.2628
month10	0.520224	1.4226828	0.365664	0.7147
month11	-1.742238	1.5319335	-1.137280	0.2558
month12	-0.957199	1.4891726	-0.642773	0.5206
DML: month2	-0.007741	0.0072987	-1.060631	0.2892
DML: month3	-0.020683	0.0058091	-3.560466	0.0004
DML: month4	-0.023433	0.0061142	-3.832590	0.0001
DML: month5	-0.028571	0.0066471	-4.298245	0.0000
DML: month6	-0.012384	0.0075008	-1.651075	0.0991
DML: month7	-0.036971	0.0053041	-6.970336	0.0000
DML: month8	-0.036261	0.0052885	-6.856600	0.0000
DML: month9	-0.002746	0.0056084	-0.489542	0.6246
DML: month10	-0.008962	0.0055068	-1.627426	0.1041
DML: month11	0.007326	0.0061465	1.191932	0.2337
DML: month12	0.007205	0.0061136	1.178596	0.2389

We covered several possible ways to model heteroskedastic data:

Type	Formula	DF	R Function
Fixed	$\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times x_i)$	0	<code>varFixed()</code>
Constant	$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma_j^2)$	j-1	<code>varIdent()</code>
Power	$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 \times x_i ^{2\delta_j})$	1 or j	<code>varPower()</code>
Exponential	$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 \times e^{2\delta_j \times x_i})$	1 or j	<code>varExp()</code>
Const.+Power	$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 \times (\delta_{1j} + x_i ^{\delta_{2j}})^2)$	1 or 2j	<code>varConstPower()</code>
Combination	Variable	Var.	<code>varComb()</code>

References

Smith, J.M., Pierce, G.J., Zuur, A.F. & Boyle, P.R. (2005). Seasonal patterns of investment in reproductive and somatic tissues in the squid *loligo forbesi*. *Aquatic Living Resources*, 18, 341–351.

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